

$$\theta = (\theta_1, \dots, \theta_k)$$

$$\hat{\theta}_{MLE} \sim ?$$

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10)

STAT 206
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In regular parametric problems, in which

$$D = Y = (Y_1, \dots, Y_n)$$

Support (Σ_i) does ^①
not depend on θ ,

$$\hat{\theta}_{MLE} \underset{RS}{\sim} N_k(\theta, \frac{1}{n} \hat{I}_{\hat{\theta}_{MLE}})$$

as $n \rightarrow \infty$
with k
fixed

in which $\hat{I}_{\hat{\theta}_{MLE}} = I_{\theta_{MLE}}^{-1}$ and

$$\hat{I}_{\hat{\theta}_{MLE}} = - \left(H \Big|_{\theta = \hat{\theta}_{MLE}} \right)$$

Hessian of log likelihood function

$$\text{bias}(\hat{\theta}_{MLE}) \rightarrow 0 \text{ as } n \rightarrow \infty$$

1 bias
ex.
for
 $\hat{\theta}_{MLE}$

$(Y_i | \mu, \sigma^2) \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$
($i = 1, \dots, n$)

②

$\theta = (\mu, \sigma^2)$

$\hat{\mu}_{MLE} = \frac{1}{n} \sum_{i=1}^n Y_i$	$E_{RS}(\hat{\mu}_{MLE}) = \mu$	$Z = (Y_1, \dots, Y_n)$ $\underline{I} = (I_1, \dots, I_n)$
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so $\hat{\mu}_{MLE}$ is unbiased for μ but standard

$\hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y})^2$	<p>estimate of σ^2 is</p> $s^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$
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s^2 is ~~preferred~~ preferred over $\hat{\sigma}_{MLE}^2$ because s^2 is unbiased for σ^2 & $\hat{\sigma}_{MLE}^2$ is not

bias($\hat{\sigma}_{MLE}^2$) = ~~$E_{RS}(\hat{\sigma}_{MLE}^2) - \sigma^2$~~ $E_{RS}(\hat{\sigma}_{MLE}^2) - \sigma^2 < 0$

$$\frac{(n-1)s^2}{n} = \frac{1}{n} \sum_{i=1}^n (Y_i - \bar{Y})^2 = \frac{n \hat{\sigma}_{MLE}^2}{n}$$

$$\begin{aligned} \text{bias}(\hat{\sigma}_{MLE}^2) &= E_{\theta, \sigma} \left(\frac{n-1}{n} s^2 \right) - \sigma^2 \quad (2) \\ &= E_{\theta, \sigma} \left[\frac{n-1}{n} E_{\theta, \sigma}(s^2) \right] - \sigma^2 \\ &= \frac{n-1}{n} \sigma^2 - \sigma^2 = \sigma^2 \left(\frac{n-1}{n} - 1 \right) \end{aligned}$$

$$= \frac{n-1}{n} \sigma^2 - \sigma^2 = \sigma^2 \left(\frac{n-1}{n} - 1 \right)$$

$$= \frac{n-1}{n} \sigma^2 - \sigma^2 = \sigma^2 \left(\frac{n-1}{n} - 1 \right)$$

$$= -\frac{\sigma^2}{n} = O\left(\frac{1}{n}\right)$$

$$\hat{\theta}_{MLE} \rightarrow \left(\hat{\mu}_{MLE}, \hat{\sigma}_{MLE}^2 \right)$$

THT 2
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$k=3$

$$\begin{aligned} \mathcal{L}_{\theta}(\tilde{x} | \tilde{x} \sim \mathcal{B}) &= n_1 \log \theta_1 \\ &+ n_2 \log \theta_2 + n_3 \log \theta_3 \end{aligned}$$

$$\tilde{\theta} = (\theta_1, \theta_2, \theta_3)$$

$$\mathcal{L}_{\theta}(\tilde{x} | \tilde{x} \sim \mathcal{B}) = n_1 \log \theta_1 + n_2 \log \theta_2 +$$

$$\theta_1 + \theta_2 + \theta_3 = 1 \rightarrow \theta_3 = 1 - \theta_1 - \theta_2 \quad n_3 \log(1 - \theta_1 - \theta_2)$$