

Mr. Fisher wants

STAT 206  
3 Feb 21

$$\hat{V}_F(\hat{\theta}_{MLE}) \rightarrow$$

wed AM  
discussion  
section/  
lecture

$$\sqrt{E_{IID}(\hat{\theta}_{MLE})} = \sqrt{\hat{V}_{IID}(\hat{\theta}_{MLE})} \quad (1)$$

$$I(\hat{\theta}_{MLE}) = \left[ -\frac{d^2}{d\theta^2} \log(\theta | \mathcal{Z} \mathcal{B}) \right]_{\theta = \hat{\theta}_{MLE}}$$

(Fisher) information provided by  
 $\hat{\theta}_{MLE}$  about  $\theta$

Mr. Fisher's recipe for (estimation-  
based) inference with parametric  
sampling models

① based on context  $P = (\Theta, \mathcal{L})$  <sup>②</sup>

consist to a single parametric sampling distribution / model  $\mathcal{M}$

$$\left( Y_i \mid \theta \in \mathcal{M} \right) \stackrel{\text{IID}}{\sim} p(y_i \mid \theta \in \mathcal{M})$$

$(i=1, \dots, n) \quad (\theta \in \Theta)$   
 $(\dim(\Theta) = 1 \text{ for now})$

i.e., we (pretend that we) know the functional form of  $p(y_i \mid \theta \in \mathcal{M})$

② The only unknown thing is  $\theta$

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③ (ex. /  $(Y_i \mid \theta \in \mathcal{B}(\text{Bernoulli sampling model})) \stackrel{\text{IID}}{\sim} \text{Bernoulli}(\theta)$   $\theta = (0, 1)$

i.e.,  $p_{Y_i}(y_i \mid \theta \in \mathcal{M}) = \theta^{y_i} (1-\theta)^{1-y_i} \mathbb{I}(y_i = \frac{0}{1})$

1  
① write down marginal sampling dist. <sup>③</sup>  
of a single  $Z_i$

$$\underline{Z} = (Z_1, \dots, Z_n)$$
$$z = (z_1, \dots, z_n)$$

② write down joint sampling dist. for  $\underline{Z}$ :  
$$P_{\underline{Z}}(z | \theta \in \mathcal{B}(\mathcal{M})) = \prod_{i=1}^n P_{Z_i}(z_i | \theta \in \mathcal{B}(\mathcal{M}))$$

③ define likelihood function:

$$l(\theta | z \in \mathcal{B}(\mathcal{M})) = c P_{\underline{Z}}(z | \theta \in \mathcal{B}(\mathcal{M}))$$

for arbitrary  $c > 0$ , & plot it

④ so  $l(\theta | z \in \mathcal{B}(\mathcal{M})) = \prod_{i=1}^n P_{Z_i}(z_i | \theta \in \mathcal{B}(\mathcal{M}))$   
for all  $z_s = \{1 \text{ or } 0\}$

ex.

$$P_{\underline{Z}}(z | \theta \in \mathcal{B}(\mathcal{M})) = \prod_{i=1}^n \theta^{z_i} (1-\theta)^{1-z_i}$$

$(s = \sum_{i=1}^n z_i)$

$$= \theta^s (1-\theta)^{n-s}$$

4 define log likelihood function: ④

$$l(\theta | \mathcal{Y} \sim \mathcal{B}(\theta)) = \log [L(\theta | \mathcal{Y} \sim \mathcal{B}(\theta))]$$

$$= \sum_{i=1}^n \log p_{\mathcal{Y}_i}(\mathcal{Y}_i | \theta \sim \mathcal{B}(\theta)) \quad \& \text{plot it}$$

ex.  $l(\theta | \mathcal{Y} \sim \mathcal{B}(\theta)) = s \log \theta + (n-s) \log(1-\theta)$   
(Bernoulli)

5 find the  $\theta$  that maximizes the log likelihood function:

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} l(\theta | \mathcal{Y} \sim \mathcal{B}(\theta))$$

(set  $\frac{d}{d\theta} l(\theta | \mathcal{Y} \sim \mathcal{B}(\theta)) = 0$  & solve for  $\theta$ )  
do this analytically, if you can;  
otherwise do it numerically

ex. (Quiz 2)

$$\frac{d}{d\theta} \ell(\theta | \mathcal{Y} \sim \mathcal{B}(n)) = 0$$

led immediately to

$$\hat{\theta}_{MLE} = \frac{\sum}{n} = \frac{s}{n} \text{ (data)}$$

random variable

$$\boxed{6} \quad \widehat{SE}(\hat{\theta}_{MLE}) = \sqrt{\widehat{V}(\hat{\theta}_{MLE})}$$

$$= \sqrt{[I(\hat{\theta}_{MLE})]^{-1}}$$

$I(\hat{\theta}_{MLE}) = O(n)$

where  $I(\hat{\theta}_{MLE}) = \left[ -\frac{d^2}{d\theta^2} \ell(\theta | \mathcal{Y} \sim \mathcal{B}(n)) \right]$

ex  $\widehat{SE}(\hat{\theta}_{MLE}) = \sqrt{\frac{\hat{\theta}_{MLE}(1-\hat{\theta}_{MLE})}{n}}$   
 $= O(\frac{1}{\sqrt{n}})$

$\theta = \hat{\theta}_{MLE}$

7] construct an approximate (large- $n$ )  
 $100(1-\alpha)\%$  likelihood-based  
confidence interval:

$$\hat{\theta}_{MLE} \pm z_{\alpha/2} \sqrt{[I(\hat{\theta}_{MLE})]^{-1}}$$

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Bayesian: nasty integrals  
(harder)

likelihood  
story: differentiation  
(easier)

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